

Health Affairs

At the Intersection of Health, Health Care and Policy

Cite this article as:

Sam Harper, Richard F. MacLehose and Jay S. Kaufman
Trends In The Black-White Life Expectancy Gap Among US States, 1990–2009
Health Affairs, 33, no.8 (2014):1375-1382

doi: 10.1377/hlthaff.2013.1273

The online version of this article, along with updated information and services, is available at:

<http://content.healthaffairs.org/content/33/8/1375.full.html>

For Reprints, Links & Permissions:

http://healthaffairs.org/1340_reprints.php

E-mail Alerts : <http://content.healthaffairs.org/subscriptions/etoc.dtl>

To Subscribe: <http://content.healthaffairs.org/subscriptions/online.shtml>

Health Affairs is published monthly by Project HOPE at 7500 Old Georgetown Road, Suite 600, Bethesda, MD 20814-6133. Copyright © 2014 by Project HOPE - The People-to-People Health Foundation. As provided by United States copyright law (Title 17, U.S. Code), no part of *Health Affairs* may be reproduced, displayed, or transmitted in any form or by any means, electronic or mechanical, including photocopying or by information storage or retrieval systems, without prior written permission from the Publisher. All rights reserved.

Not for commercial use or unauthorized distribution

By Sam Harper, Richard F. MacLehose, and Jay S. Kaufman

Trends In The Black-White Life Expectancy Gap Among US States, 1990–2009

DOI: 10.1377/hlthaff.2013.1273
HEALTH AFFAIRS 33,
NO. 8 (2014): 1375–1382
©2014 Project HOPE—
The People-to-People Health
Foundation, Inc.

ABSTRACT Nationwide differences in US life expectancy trends for blacks and whites may mask considerable differences by state that are relevant to policies aimed at reducing health inequalities. We calculated annual state-specific life expectancies for blacks and whites from 1990 to 2009 using age-specific mortality counts and census-based denominators. Nationally, the black-white difference in life expectancy at birth shrank during the period by 2.7 years for males (from 8.1 to 5.4 years) and by 1.7 years for females (from 5.5 to 3.8 years). We found considerable variation across states in both the magnitude of the life expectancy gap (approximately fifteen years) and the change during the past two decades (about six years). Decomposition analysis showed that New York made the most profound contribution to reducing the gap, but less favorable trends in a number of states, notably California and Texas, kept the gap from shrinking further. Large state variations in the pace of change in the racial gap in life expectancy suggest that state-specific determinants merit further investigation.

Sam Harper (sam.harper@mcgill.ca) is an assistant professor of epidemiology, biostatistics, and occupational health at McGill University, in Montreal, Quebec.

Richard F. MacLehose is an associate professor of epidemiology and community health at the University of Minnesota, in Minneapolis.

Jay S. Kaufman is a professor of epidemiology, biostatistics, and occupational health at McGill University.

Life expectancy continues to rise in the United States for all major population groups.¹ However, the aggregate data obscure important variations at the subnational level.^{2–5}

Racial and ethnic inequalities in life expectancy, especially those between blacks and whites, remain a vital issue for policy makers. This is not only because inequalities reflect the cumulative effects of disadvantage in disease risk factors and medical interventions across the life course,^{6,7} but also because there is near-universal agreement that such inequalities are preventable.⁸ The national black-white gap in life expectancy at birth continues to narrow over time.^{9,10} Nonetheless, the gap's persistence represents inadequate progress toward social justice in health.¹¹

African American populations are highly concentrated across states. The 2010 census revealed that in twenty-one of the fifty states the black population is at least 10 percent of the

total. Mississippi has the largest proportion of blacks (37 percent), and Connecticut has the smallest (10 percent).¹² The census also reported that in eighteen states the black population is greater than 1 million. Of these, New York has the largest population of blacks (3.1 million), and Tennessee has the smallest (1.1 million).

The national life expectancy gap between blacks and whites at any given time is a function of the individual state-specific gaps and the racial distribution across states. Similarly, changes in the overall national life expectancy gap result from the combined effect of each of the state-specific changes in the gap and changes in the racial distribution across states. As a result, the national black-white gap could be shrinking because of stronger growth in life expectancy among blacks in each state or because of migration to states with smaller black-white life expectancy inequalities.^{13,14}

Moreover, much of American social and health

policy is set at the state level, and there is wide variation across states in health systems and forms of economic intervention. Thus, understanding state-specific trajectories is an essential step toward explaining the patterns seen in the aggregate trends.

The estimation of state-specific inequality in black-white life expectancy is hampered by the number of states with very small black populations and, thus, very small numbers of age-specific black deaths. This problem necessitates some kind of smoothing procedure in order to obtain reasonable estimates in cases where the data are sparse.¹⁵

Narzleen Bharmal and coauthors evaluated state-specific racial differences in life expectancy by pooling several years of data (1997–2004), but they excluded eleven states that had too few events for stable models.⁵ Such pooling could mask important trends. For example, between the late 1990s and late 2000s the United States experienced rising rates of deaths from unintentional drug poisoning—trends that affected whites to a greater extent than blacks.¹⁶ If these trends also differed by state—and there is strong evidence of state variation in drug poisoning mortality¹⁶—pooling mortality data could hide important state variations in the black-white life expectancy gap.

In this article we used data from a broader range of years (1990–2009) to estimate annual state-specific inequalities in life expectancy at birth between blacks and whites for all fifty states and the District of Columbia. We estimated these trajectories separately for males and females because previous work has indicated that the causes of changes in the black-white gap differ for males and females.^{9,10,17} Our estimates allow an understanding of temporal changes, an assessment of how the racial gap has evolved over time from state to state, and an assessment of which states have contributed most to changes in the racial gap during this twenty-year period.

We used Bayesian smoothing methods that stabilized rates and allowed the calculation of 95 percent credible interval estimates (a credible interval is the Bayesian analogue to a traditional confidence interval) for categories with small black populations. This allowed us to make the first comprehensive depiction of changing state-level racial inequalities in life expectancy over an extended period of time, with important implications for national trends.

Study Data And Methods

DATA The data source for this study was the National Vital Statistics System. The system, which is maintained by the National Center for Health

Statistics (NCHS), collects information on all deaths occurring in the United States each year. For the fifty states and the District of Columbia, we calculated state-specific all-cause mortality rates for infants, children ages 1–4, and the remaining five-year age categories (up to ages 85 and older) from 1990 to 2009 for black and white males and females. We extracted the data using software from the National Cancer Institute’s Surveillance, Epidemiology, and End Results Program,¹⁸ for which the NCHS provides the underlying mortality data and population estimates.¹⁹

We did not restrict our analysis to non-Hispanics because data on mortality by ethnicity are not considered to be reliable for all states during this period.²⁰ However, we conducted a sensitivity analysis for states reported to have reliable mortality data by ethnicity.²¹ Five states (Idaho, Montana, North Dakota, South Dakota, and Vermont) with very small black populations were excluded from the analysis (approximately 0.1 percent of the total US black population).¹² Results are therefore shown for the remaining forty-six states (including the District of Columbia), which we refer to as the “national” sample.

We used data for nineteen age categories,²² forty-six state entities, two races, two sexes, and twenty years. Thus, there were 69,920 crude mortality rates that could be calculated from the data. However, given the small black populations in many states, these crude estimates suffer from a lack of precision. We used a Bayesian extension of standard life-table methods to estimate life expectancy at birth in each state and year for four race-sex groups: black females, white females, black males, and white males.

STATISTICAL METHODS To calculate life expectancies, we obtained denominator counts from the Census Bureau.²³ We modeled the number of deaths in each state-age-sex-race category in each year using Poisson regression, with the size of each group as an offset term in the regression.

The NCHS reports zero counts for deaths when, for a given group (for example, black males ages 15–19 in Wyoming in 2001), no deaths have been reported. However, for reasons of confidentiality, the NCHS suppresses the number of deaths if the count is 1–9.²⁴ For these categories (12.9 percent of all 69,920 state-age-sex-race observations), we imputed the unobserved count using a truncated Poisson regression to accommodate the fact that the true number of deaths was censored (for the complete specification of the model, see the online Appendix).²⁵

Furthermore, because mortality rates were less stable in categories that had a small number of deaths (for example, in places with relatively

The surveillance of average life expectancy trends and social inequalities is a prime public health objective.

small populations or where death rates were very low), we opted to smooth all state-age-sex-race death rates across years using a conditional autoregressive prior specification. This specification allowed the death rates for each state-age-sex-race group to be smoothed toward the rate in adjacent years. For example, the death rate for black men ages 35–39 in New Hampshire in 2003 was smoothed using data for those in 2002 and 2004 to provide a more stable estimate.

We then used Markov chain Monte Carlo methods to obtain state-specific annual mortality rates for each race by sex and age. We converted the mortality rates from this model to risks and then used standard demographic methods to calculate life expectancy at birth.²⁶ Specifically, we used 10,000 samples from the posterior distribution of each category-specific rate to compute life expectancies. We used the mean of these 10,000 samples as the point estimate for life expectancy for blacks and whites in each state, and we used the 2.5th and 97.5th percentiles of the distribution to estimate precision.

To assess regional variation, we combined the estimates for the change in the gap in each state using an inverse-variance weighted fixed effects meta-analysis. We tested for heterogeneity within and across census divisions using Cochran's *Q* statistic.²⁷

LIMITATIONS Several limitations to our study are worth noting. We did not restrict our sample to non-Hispanics because of data reliability concerns. Thus, our results might not necessarily be applicable to non-Hispanic populations. However, our sensitivity analysis for states considered to have reliable reporting of ethnicity on death certificates was consistent with the estimates for Hispanics and non-Hispanics combined. If anything, we may have underestimated the magnitude of the decline in the life expectancy gap between non-Hispanic blacks and non-Hispanic whites because increasing numbers of Hispanics have contributed to greater life

expectancy improvement among whites than among blacks.

To stabilize mortality rates, we implemented Bayesian models that borrowed information on mortality rates for each state-age-sex-race category from adjacent years. As a result, estimated changes over time could be less pronounced and thus more conservative. However, there is evidence that the use of Bayesian methods for life expectancy estimation in smaller samples leads to improvements over more traditional life-table methods.²⁸

Study Results

Point estimates and 95 percent credible intervals (CIs) for black and white life expectancy at birth and the black-white life expectancy gap for each state in 1990 and 2009 are given in Appendix Exhibits 1 and 2.²⁵ The trend in the black-white gap and its CI for each state are shown in Appendix Exhibits 3 and 4²⁵ for males and females, respectively.²⁹

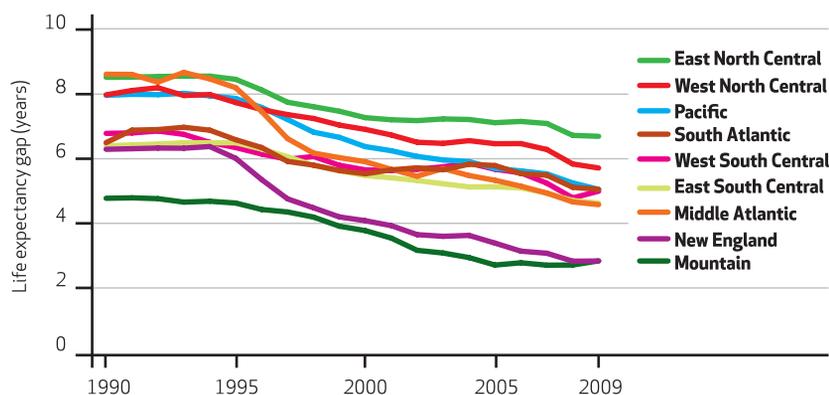
During the twenty-year study period, the national life expectancy gap (weighting each state by its total black and white population) between blacks and whites shrank by 2.7 years for males (from 8.1 years in 1990 to 5.4 years in 2009; Appendix Exhibit 1), and by 1.7 years for females (from 5.5 years to 3.8 years; Appendix Exhibit 2). The estimated state-specific racial gap in life expectancy among males in 1990 ranged from 14.4 years (95% CI: 13.3, 15.4) in the District of Columbia to 0.0 years (95% CI: -5.2, 5.3) in New Hampshire. Both Alaska and Hawaii actually showed a black advantage for males in 1990, but these gaps were very imprecisely estimated (Appendix Exhibit 1).²⁵ For females, the District of Columbia and New Hampshire also had the largest (10.4 years; 95% CI: 9.5, 11.2) and smallest (1.2 years; 95% CI: -3.9, 6.2) gaps, respectively, in 1990.

In 2009, despite overall improvement at the national level, the enormous racial gap in the District of Columbia remained almost completely unchanged for both males (+0.4 years; 95% CI: -1.1, 1.8) and females (+0.2 years; 95% CI: -1.1, 1.6). Life expectancy there remained dramatically more unequal than in every other state.

The largest decrease in the gap for males occurred in New York (-5.6 years; 95% CI: -6.0, -5.1), largely because of a dramatic increase in black life expectancy, from 63.9 years to 75.4 years. This was a gain of 11.5 years; the national average gain was 6.5 years. For females, the biggest improvement in the gap was observed in Wyoming (-4.1 years), but it was estimated with substantial imprecision (95% CI: -11.8, 1.3). Stronger and more reliably estimated

EXHIBIT 1

Trends In The Black-White Life Expectancy Gap For Males, By Census Region, 1990–2009



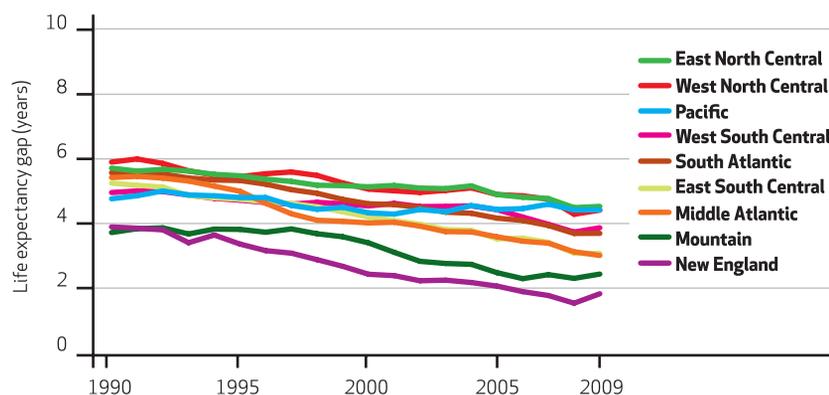
SOURCE Authors' analysis of data from the National Vital Statistics System and the Census Bureau.
NOTE Regional estimates were calculated as weighted averages, with each state weighted by the inverse of the variance of its black-white life expectancy gap in each year.

gains among females were seen in New Jersey, New York, Oregon, and Rhode Island.

Exhibits 1 and 2 show estimates of trends in the black-white life expectancy gap in each of the nine census divisions for males and females, respectively (for a map of the divisions, see Appendix Exhibit 5).²⁵ These trends illustrate the considerable variation in both levels and trends in the gap (the state-specific trends for both sexes are shown in Appendix Exhibits 6 and 7,²⁵ respectively). The Middle Atlantic and New England areas showed stronger declines in the black-white life expectancy gap for males during the mid- to late 1990s, compared to other areas. The pace of decline in the gap for females was more similar across areas, except the Pacific di-

EXHIBIT 2

Trends In The Black-White Life Expectancy Gap For Females, By Census Region, 1990–2009



SOURCE Authors' analysis of data from the National Vital Statistics System and the Census Bureau.
NOTE Regional estimates were calculated as weighted averages, with each state weighted by the inverse of the variance of its black-white life expectancy gap in each year.

vision, where the trend was essentially flat.

Exhibit 3 shows point estimates and measures of precision for the change in the black-white gap in each census division, along with heterogeneity statistics (the complete forest plots and additional statistical information are given in Appendix Exhibits 8, 8A, and 9).²⁵ New England and the Middle Atlantic showed greater-than-average declines in the gap for males (approximately 3–4 years), whereas declines were smaller than average in the East and West South Central (less than 2 years). For females the pattern was largely similar to that for males, with large improvements coming in the Middle Atlantic and New England, but it was dissimilar in that the East South Central region showed a larger-than-average decline in the gap.

For males we found evidence of heterogeneity in the gap among states within every division (all $p \leq 0.01$), with the exception of the East South Central (Cochran's $Q = 2.92$; degrees of freedom = 3; $p = 0.40$) and Mountain ($Q = 4.20$; $df = 5$; $p = 0.51$) areas (Exhibit 3). For females we also found considerable heterogeneity within most census divisions.

Among the states with large black populations, New York was notable for its particularly dramatic decline in the gap for both sexes. This is in stark contrast to California, which began the period with roughly the same magnitude of inequality as New York but which maintained a roughly constant gap, decreasing by only about a year for males and not at all for females. However, this was not because California showed little improvement in life expectancy. Blacks and whites showed similar above-average increases in life expectancy in California (about 3.5 years for females and 5.0 years for males; see Appendix Exhibits 10 and 11).²⁵ This reinforces the fact that the changes in the black-white life expectancy gap may arise from various combinations of changes in the component rates. Other relatively large states with poorer improvements in the black-white gap for males included Ohio and Wisconsin.

The pattern for females was somewhat different, but here again New York was the best-performing large state. California registered virtually no change in the life expectancy gap among females during the study period. One state, Wisconsin, had a significantly larger gap for females in 2009 than it did in 1990 (an increase of 1.57 years; 95% CI: 0.64, 2.47).

A simple decomposition analysis of the change in the national gap (described in detail in the Appendix)²⁵ found that for males and females, respectively, 96 percent (2.6 years) and 98 percent (1.6 years) of the decrease in the black-white life expectancy gap was attributable to differen-

EXHIBIT 3
Changes In The Black-White Life Expectancy Gap, By Census Division, 1990–2009

| Sex and region | Change in the black-white life expectancy gap | | | Tests of heterogeneity within census division | |
|--------------------|---|----------------|--------|---|----------------------------|
| | Years | 95% CI | Weight | p value | % of variance ^a |
| MALES | | | | | |
| New England | -3.32 | (-3.97, -2.66) | 2.46 | 0.01 | 69.70 |
| Middle Atlantic | -4.20 | (-4.49, -3.91) | 13.05 | <0.01 | 97.60 |
| East North Central | -1.85 | (-2.10, -1.59) | 16.79 | <0.01 | 94.20 |
| West North Central | -2.22 | (-2.72, -1.72) | 4.36 | <0.01 | 86.10 |
| South Atlantic | -2.87 | (-3.06, -2.67) | 28.58 | <0.01 | 89.30 |
| East South Central | -1.80 | (-2.10, -1.49) | 11.56 | 0.40 | 0.00 |
| West South Central | -1.76 | (-2.04, -1.47) | 13.33 | <0.01 | 82.90 |
| Mountain | -1.98 | (-2.67, -1.29) | 2.25 | 0.51 | 0.00 |
| Pacific | -1.37 | (-1.75, -1.00) | 7.62 | <0.01 | 76.30 |
| Overall | -2.45 | (-2.55, -2.34) | 100.00 | <0.01 | 92.20 |
| FEMALES | | | | | |
| New England | -2.04 | (-2.68, -1.40) | 2.36 | 0.58 | 0.00 |
| Middle Atlantic | -2.46 | (-2.73, -2.19) | 13.50 | <0.01 | 94.80 |
| East North Central | -1.16 | (-1.40, -0.93) | 17.30 | <0.01 | 94.30 |
| West North Central | -1.41 | (-1.87, -0.95) | 4.59 | <0.01 | 79.40 |
| South Atlantic | -1.76 | (-1.95, -1.58) | 27.80 | <0.01 | 82.00 |
| East South Central | -2.11 | (-2.40, -1.82) | 11.72 | 0.01 | 73.40 |
| West South Central | -1.07 | (-1.34, -0.80) | 13.37 | 0.22 | 32.60 |
| Mountain | -1.21 | (-1.91, -0.51) | 1.98 | 0.88 | 0.00 |
| Pacific | -0.33 | (-0.69, 0.04) | 7.38 | 0.05 | 57.10 |
| Overall | -1.57 | (-1.67, -1.48) | 100.00 | <0.01 | 86.60 |

SOURCE Authors' analysis of data from the US National Vital Statistics System and the US Census Bureau. **NOTES** Summary estimates were calculated using inverse-variance weighted fixed effects meta-analysis. States with more precise estimates were given greater weight. CI is confidence interval. Note that in the text, "CI" refers to "credible interval." ^aWithin each division attributable to heterogeneity across states (I^2).

tial changes in life expectancy instead of differential migration. This suggests that migration played almost no role in decreasing the national black-white life expectancy gap.

Because faster mortality declines among blacks were the chief reason for the decline in the national life expectancy gap, we present below only the results of our state-specific decomposition for counterfactual mortality changes. Appendix Exhibits 12 and 13²⁵ show, for males and females, respectively, the contribution of each state to the change in the national black-white gap from 1990 to 2009.

States are rank-ordered on the vertical axis by their observed twenty-year change in black life expectancy. When we weighted states by their total black and white populations, the observed national change in the black-white life expectancy gap was -2.7 years for males and -1.6 years for females, and the plotted point for each state along the horizontal axis is the counterfactual national change in the gap that would have been observed had blacks and whites in that state experienced the national average trends in life expectancy growth.

New York was a clear outlier for both males

and females. For example, Appendix Exhibit 12²⁵ shows a value of -2.5 for New York, indicating that if it had experienced the national average life expectancy growth for whites (0.25 percent per year) and blacks (0.48 percent per year) instead of its actual growth rates (0.40 percent and 0.83 percent, respectively), the national black-white gap for males would have declined by only 2.5 years instead of 2.7 years.

In contrast, the observed life expectancy growth rates for white and black males in California were 0.35 percent and 0.46 percent, respectively. The faster-than-average white life expectancy growth coupled with just average black life expectancy growth is why California kept the national black-white gap from declining even further than it did.

The extreme contributions of New York and California are readily evident for both sexes. To a lesser extent than New York, Florida, Illinois, and New Jersey also played important roles in reducing the black-white gap. In contrast, Alabama, Louisiana, and Mississippi kept the national black-white gap from falling by more than it would have otherwise.

A sensitivity analysis of state trends in the

black-white gap restricted to non-Hispanics largely confirmed the results for Hispanics and non-Hispanics combined (for estimates for states considered to have reliable reporting in 1990 and 2009, see Appendix Exhibits 14 and 15).²⁵ Restricting the sample to non-Hispanics led to a slightly larger reduction in the overall gap (2.9 years for males compared to 2.7 years for all ethnicities, and 1.8 versus 1.7 years for females; Appendix Exhibits 13 and 14, compared to Appendix Exhibits 1 and 2).²⁵ This is primarily a result of smaller improvements in life expectancy for whites when Hispanics were excluded.

Discussion

Racial differences in life expectancy can be difficult to interpret because the same inequality may arise from a multitude of black and white life expectancies. The study of changes in inequalities over time is even more challenging, since the gap can increase when both groups experience rising life expectancy, and it can decrease even if both groups experience falling life expectancy. For example, both black and white males in the District of Columbia gained nearly ten years of life expectancy from 1990 to 2009, yet the life expectancy gap there remained constant at nearly fifteen years and easily remains the largest among the states.

Nonetheless, the surveillance of average life expectancy trends and social inequalities is a prime public health objective³⁰ and is necessary for the assessment of progress toward health equality. Our analysis provides a high-resolution picture of the evolution of black-white racial mortality inequalities during the past two decades. This is the first article to detail these values by state and year, with formal assessments of imprecision. Our results depict considerable heterogeneity across states in their trajectories of improvement and raise many important questions about the origins of these patterns in state policies, migration patterns, and the geographic patterns of important risk factors.^{31–33}

For instance, New York City has experienced substantial improvements in homicide³⁴ and HIV/AIDS mortality during the past two decades.³⁵ These trends have disproportionately benefited blacks in absolute terms and may go some way toward explaining the much larger change in black life expectancy in New York. Samuel Preston and Irma Elo attribute these improvements largely to local policies, particularly to the aggressive identification of and treatment for people with HIV/AIDS.³⁶ They also provide evidence that reductions in drug- and alcohol-related deaths contributed substantially to New York City's large rise in life expectancy during

the past two decades.

The pattern in New York is in stark contrast to that in California, which experienced notably strong declines in tobacco use during past decades³⁷ and has outpaced the nation in declines in heart disease and smoking-attributable mortality.^{38,39} These gains may have differentially benefited whites relative to blacks when compared to progress in other states. Detailed trends on cause-specific mortality will help shed light on this question.

There was no area in which males or females experienced a decrease in life expectancy. However, improvements were sufficiently uneven that the black-white gap worsened in some places (notably among females in Wisconsin), even as mortality rates were falling. The explanations for these divergent patterns rest on a myriad of social, economic, and demographic shifts, in addition to policy differences and the legacy of historical inequalities that have not been quickly resolved. For example, the District of Columbia continues to have the largest magnitude of racial difference in life expectancy of any area, despite the fact that the District's population experienced a profound demographic shift during the study period—from being nearly 70 percent black in 1990 to being only about 50 percent black in 2009.

The diversity of state-specific trajectories in the black-white life expectancy gap suggests important avenues for investigating potential causes of and solutions to racial inequalities in health. Considerable previous work has focused on the lifetime exposure of blacks and whites to substantially different economic and environmental conditions, largely structured through residential segregation.^{40,41} These factors should be investigated at the state level, but national socioeconomic trends do not correlate well with the continued declines in the black-white life expectancy gap.

For example, trends in the black-white income and wealth gaps during the study period were generally unfavorable—particularly for black males, who experienced the strongest life expectancy gains—and there is little evidence of black progress in homeownership or full employment.^{42,43}

It is also difficult to reconcile patterns in residential segregation with our results. There were larger decreases in segregation in the Mountain and Pacific regions, where we found smaller (though less precise) decreases in the black-white life expectancy gap. However, there has been some black-white convergence in educational attainment. Recent estimates of widening life expectancy differences by education⁴⁴ suggest that state differences in black-white educa-

Future work should probe deeper into these patterns by examining the age-specific mortality rates and primary causes of death.

tion inequalities could play some role in explaining divergent state trends.

Race-specific state trends in environmental risk factors for major causes of mortality are likely avenues for further investigation, but few studies have assessed state trends by race. A recent study⁴⁵ on state trends from 2001 to 2010 in preventable cardiovascular mortality found that all states showed declines. However, progress was slowest across Appalachia and in much of the East and West South Central regions. These areas were where we found slower progress in reducing the black-white gap. They were also among the areas that performed worst in increasing life expectancy for both blacks and whites. This may suggest that inadequate progress in reducing cardiovascular risk factors and access to treatment are negatively affecting both races.

The strengths of our study include the use of complete population data, which eliminated the need for sampling, and the use of statistical smoothing to increase stability and allow for a greater number of estimates than have been reported previously. We also accounted statistically for the censored values of small counts and reported uncertainty intervals for the change in life expectancy gaps over time.

Future work should probe deeper into these patterns by examining the age-specific mortality rates and primary causes of death. The end of the period of high HIV/AIDS mortality occurred during our study period. In addition, homicide rates have declined substantially in many areas. Both of these changes are known to make important contributions to the narrowing of the racial gap.¹⁰ State-level policies regarding policing, sentencing, and laws related to illicit drugs have also changed importantly during our observation period. This has led to dramatic shifts in incarceration rates, which may also have important impacts on mortality-rate inequalities.⁴⁶

A detailed examination of these patterns by state and time has not yet been conducted, but it should be an obvious next step. This must naturally involve the intense scrutiny of states that have remarkable patterns of change, such as California, the District of Columbia, New York, and Wisconsin.

The heterogeneous gains across southern states should also be a target of further inquiry. These areas include large numbers of blacks and whites, make important contributions to the national statistics, and have unique stories to tell. ■

This work was funded by a Chercheur Boursier Junior 2 award from the Fonds de la Recherche en Santé du Québec and the Canada Research Chairs

Program. No sponsors had any input into the research. The raw data and statistical code used to generate the results in this study are available with

unrestricted open access from the corresponding author's Dataverse (<http://dvn.iq.harvard.edu/dvn/dv/samharper>).

NOTES

- Centers for Disease Control and Prevention. QuickStats: life expectancy at birth, by race and sex—United States, 2000–2009. *MMWR Morb Mortal Wly Rep*. 2011; 60(18):588.
- Ezzati M, Friedman AB, Kulkarni SC, Murray CJL. The reversal of fortunes: trends in county mortality and cross-county mortality disparities in the United States. *PLoS Med*. 2008;5(5):e119.
- Kindig DA, Cheng ER. Even as mortality fell in most US counties, female mortality nonetheless rose in 42.8 percent of counties from 1992 to 2006. *Health Aff (Millwood)*. 2013;32(3):451–8.
- Kulkarni SC, Levin-Rector A, Ezzati M, Murray CJ. Falling behind: life expectancy in US counties from 2000 to 2007 in an international context. *Popul Health Metr*. 2011; 9(1):16.
- Bharmal N, Tseng CH, Kaplan R, Wong MD. State-level variations in racial disparities in life expectancy. *Health Serv Res*. 2012;47(1 Pt 2): 544–55.
- Hummer RA, Chinn JJ. Race/ethnicity and U.S. adult mortality: progress, prospects, and new analyses. *Du Bois Rev*. 2011;8(1):5–24.
- Smedley BD, Stith AY, Nelson AR, editors. *Unequal treatment: confronting racial and ethnic disparities in health care*. Washington (DC): National Academies Press; 2003.
- Koh HK, Graham G, Glied SA. Reducing racial and ethnic disparities: the action plan from the Department of Health and Human Services. *Health Aff (Millwood)*. 2011;30(10): 1822–9.
- Harper S, Rushani D, Kaufman JS. Trends in the black-white life expectancy gap, 2003–2008. *JAMA*. 2012;307(21):2257–9.
- Harper S, Lynch J, Burris S, Davey Smith G. Trends in the black-white life expectancy gap in the United States, 1983–2003. *JAMA*. 2007; 297(11):1224–32.
- Satcher D, Fryer GE Jr, McCann J,

- Troutman A, Woolf SH, Rust G. What if we were equal? A comparison of the black-white mortality gap in 1960 and 2000. *Health Aff (Millwood)*. 2005;24(2):459–64.
- 12 Census Bureau. 2010 Census interactive population search [Internet]. Washington (DC): Census Bureau; [cited 2014 Jun 30]. Available from: <http://www.census.gov/2010census/popmap/ipmtext.php>
 - 13 Hunt MO, Hunt LL, Falk WW. Twenty-first-century trends in black migration to the U.S. South: demographic and subjective predictors. *Soc Sci Q*. 2013;94(5):1398–1413.
 - 14 Hunt LL, Hunt MO, Falk WW. Who is headed south? U.S. migration trends in black and white, 1970–2000. *Soc Forces*. 2008;87(1):95–119.
 - 15 Fahrmeir L, Kneib T. Bayesian smoothing and regression for longitudinal, spatial and event history data. Oxford: Oxford University Press; 2011.
 - 16 Warner M, Chen LH, Makuc DM, Anderson RN, Miniño AM. Drug poisoning deaths in the United States, 1980–2008 [Internet]. Hyattsville (MD): National Center for Health Statistics; 2011 Dec [cited 2014 Jun 20]. (NCHS Data Brief No 81). Available from: <http://www.cdc.gov/nchs/data/databriefs/db81.pdf>
 - 17 Elo I, Drevenstedt GL. Cause-specific contributions to sex differences in adult mortality among whites and African Americans between 1960 and 1995. *Demogr Res*. 2005;13(19):485–520.
 - 18 National Cancer Institute. Surveillance, Epidemiology, and End Results Program: SEER*Stat software: latest release: version 8.1.5 [Internet]. Bethesda (MD): NCI; 2013 Mar 31 [cited 2014 Jun 20]. Available from: <http://www.seer.cancer.gov/seerstat>
 - 19 National Cancer Institute. Surveillance, Epidemiology, and End Results Program: US mortality data [Internet]. Bethesda (MD): NCI; [cited 2014 Jun 20]. Available from: <http://www.seer.cancer.gov/mortality/>
 - 20 National Cancer Institute. Surveillance, Epidemiology, and End Results Program: policy for calculating Hispanic mortality for 1990+ data [Internet]. Bethesda (MD): NCI; [cited 2014 Jun 20]. Available from: http://seer.cancer.gov/seerstat/variables/mort/origin_recode_1990+/
 - 21 States in which data were considered unreliable by the National Cancer Institute for any year between 1990 and 2009 were Connecticut, the District of Columbia, Louisiana, Maine, Maryland, Minnesota, Mississippi, New Hampshire, New York, North Dakota, Oklahoma, South Carolina, Vermont, and Virginia. For specific years, see National Cancer Institute. Surveillance, Epidemiology, and End Results Program: mortality (all causes) Hispanic index [Internet]. Bethesda (MD): NCI; [cited 2014 Jun 20]. Available from: http://seer.cancer.gov/seerstat/variables/mort/origin_recode_1990+yr1969_2010/mort.index_1990to2010.pdf
 - 22 The categories were < 1 year and 1–4, 5–9, 10–14, 15–19, 20–24, 25–29, 30–34, 35–39, 40–44, 45–49, 50–54, 55–59, 60–64, 65–69, 70–74, 75–79, 80–84, and 85+ years.
 - 23 Centers for Disease Control and Prevention. U.S. census populations with bridged race categories [Internet]. Atlanta (GA): CDC; [last updated 2014 Jun 6] cited 2014 Jun 20]. Available from: http://www.cdc.gov/nchs/nvss/bridged_race.htm
 - 24 Centers for Disease Control and Prevention. CDC WONDER: data use restrictions [Internet]. Atlanta (GA): CDC; [last reviewed 2014 May 30; cited 2014 Jun 30]. Available from: <http://wonder.cdc.gov/datause.html>
 - 25 To access the Appendix, click on the Appendix link in the box to the right of the article online.
 - 26 Preston SH, Heuveline P, Guillot M. Demography: measuring and modeling population processes. Malden (MA): Blackwell Publishers; 2001.
 - 27 Higgins JP, Thompson SG. Quantifying heterogeneity in a meta-analysis. *Stat Med*. 2002;21(11):1539–58.
 - 28 Jonker MF, van Lenthe FJ, Congdon PD, Donkers B, Burdorf A, Mackenbach JP. Comparison of Bayesian random-effects and traditional life expectancy estimations in small-area applications. *Am J Epidemiol*. 2012;176(10):929–37.
 - 29 Appendix Exhibits 10 and 11 (see Note 25) also show the absolute change in black and white life expectancy and 95% credible intervals for males and females, respectively.
 - 30 Koh HK. A 2020 vision for healthy people. *N Engl J Med*. 2010;362(18):1653–6.
 - 31 Murray CJL, Kulkarni SC, Michaud C, Tomijima N, Bulzacchelli MT, Iandiorio TJ, et al. Eight Americas: investigating mortality disparities across races, counties, and race-counties in the United States. *PLoS Med*. 2006;3(9):e260.
 - 32 Danaei G, Rimm EB, Oza S, Kulkarni SC, Murray CJL, Ezzati M. The promise of prevention: the effects of four preventable risk factors on national life expectancy and life expectancy disparities by race and county in the United States. *PLoS Med*. 2010;7(3):e1000248.
 - 33 Wang H, Schumacher AE, Levitz CE, Mokdad AH, Murray CJL. Left behind: widening disparities for males and females in US county life expectancy, 1985–2010. *Popul Health Metr*. 2013;11(1):8.
 - 34 Baumer EP, Wolff KT. Evaluating contemporary crime drop(s) in America, New York City, and many other places. *Justice Q*. 2014;31(1):5–38.
 - 35 Chiasson MA, Berenson L, Li W, Schwartz S, Singh T, Forlenza S, et al. Declining HIV/AIDS mortality in New York City. *J Acquir Immune Defic Syndr*. 1999;21(1):59–64.
 - 36 Preston SH, Elo IT. Anatomy of a municipal triumph: New York City's upsurge in life expectancy. *Popul Dev Rev*. 2014;40(1):1–29.
 - 37 Levy DT, Hyland A, Higbee C, Remer L, Compton C. The role of public policies in reducing smoking prevalence in California: results from the California tobacco policy simulation model. *Health Policy*. 2007;82(2):167–85.
 - 38 Cowling DW, Yang J. Smoking-attributable cancer mortality in California, 1979–2005. *Tob Control*. 2010;19(Suppl 1):i62–7.
 - 39 Fichtenberg CM, Glantz SA. Association of the California Tobacco Control Program with declines in cigarette consumption and mortality from heart disease. *N Engl J Med*. 2000;343(24):1772–7.
 - 40 Williams DR, Jackson PB. Social sources of racial disparities in health. *Health Aff (Millwood)*. 2005;24(2):325–34.
 - 41 Geronimus AT, Thompson JP. To denigrate, ignore, or disrupt: racial inequality in health and the impact of a policy-induced breakdown of African American communities. *Du Bois Rev*. 2004;1(2):247–79.
 - 42 Hirsch BT, Winters JV. An anatomy of racial and ethnic trends in male earnings in the U.S. *Rev Income Wealth*. 2013 Aug 22. [E-pub ahead of print].
 - 43 Pew Research Center. King's dream remains an elusive goal; many Americans see racial disparities [Internet]. Washington (DC): The Center; 2013 Aug 22 [cited 2014 Jun 20]. Available from: http://www.pewsocialtrends.org/files/2013/08/final_full_report_racial_disparities.pdf
 - 44 Olshansky SJ, Antonucci T, Berkman L, Binstock RH, Boersch-Supan A, Cacioppo JT, et al. Differences in life expectancy due to race and educational differences are widening, and many may not catch up. *Health Aff (Millwood)*. 2012;31(8):1803–13.
 - 45 Centers for Disease Control and Prevention. Vital signs: avoidable deaths from heart disease, stroke, and hypertensive disease—United States, 2001–2010. *MMWR Morb Mortal Wkly Rep*. 2013;62(35):721–27.
 - 46 Wildeman C. Imprisonment and (inequality in) population health. *Soc Sci Res*. 2012;41(1):74–91.